

# Lecture 17: Moving Beyond Linearity/Nonparametric Regression

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# Moving Beyond Linearity

The linearity assumption is almost always an approximation, and sometimes a poor one.

We can improve upon least squares using regularization  $\rightarrow$  reducing the complexity of the linear model. But we are still using a linear model.

We consider the following extensions to relax the linearity assumption.

- Polynomial regression
- Step functions
- Regression splines
- Smoothing splines
- Local regression
- Generalized additive models

# Polynomial Regression

- The **polynomial regression**

$$y_i = \beta_0 + \beta_1 x_i + \beta_2 x_i^2 + \dots + \beta_d x_i^d + \epsilon_i,$$

where  $\epsilon_i$  is the error term.

- The coefficients can be estimated using least squares linear regression.
- Not really interested in the coefficients; more interested in the fitted function values at any value  $x_0$ :

$$\hat{f}(x_0) = \hat{\beta}_0 + \hat{\beta}_1 x_0 + \hat{\beta}_2 x_0^2 + \dots + \hat{\beta}_d x_0^d.$$

- There is a simple formula to calculate the pointwise standard error of  $\hat{f}(x_0)$ . The pointwise confidence interval is  $\hat{f}(x_0) \pm 2 \cdot se[\hat{f}(x_0)]$ .
- We either fix the degree  $d$  at some reasonably low value ( $\leq 3$  or  $4$ ), else use cross-validation to choose  $d$ .

# Polynomial Regression

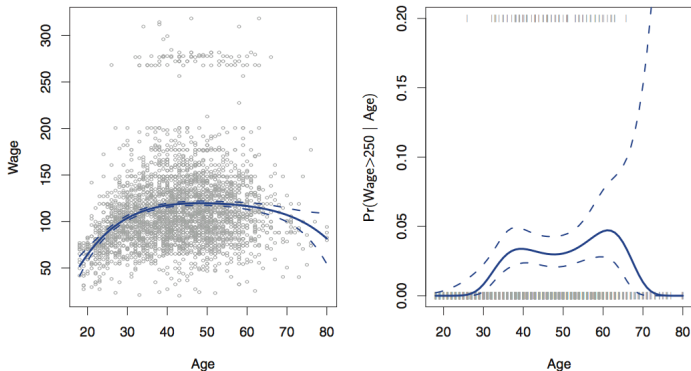
- The polynomial regression can be used for logistic regression

$$\text{logit } P(y_i = 1|x_i) = \beta_0 + \beta_1 x_i + \beta_2 x_i^2 + \dots + \beta_d x_i^d.$$

- Can do separately on several variables (see GAMs later).

# The Wage Data

## Degree-4 Polynomial



Left: The solid blue curve is a degree-4 polynomial of wage as a function of age, fit by least squares. The dotted curves indicate an estimated 95 % confidence interval. Right: We model the binary event  $\text{wage} > 250$  using logistic regression, with a degree-4 polynomial.

# Step Functions

- The polynomial regression imposes a global structure on the non-linear function of  $X$ .
- The **step function** approach avoids such a global structure. Here we break the range of  $X$  into bins, and fit a different constant in each bin. Define

$$C_0(X) = I(X < c_1), \quad C_1(X) = I(c_1 \leq X < c_2), \dots, \quad C_K(X) = I(c_K \leq X),$$

where  $c_1, c_2, \dots, c_K$  are  $K$  cutpoints in the range of  $X$ . Basically,  $C_0(X), \dots, C_K(X)$  are  $K + 1$  dummy variables, and the summation is 1.

- We then use least squares to fit a linear model using  $C_1(X), C_2(X), \dots, C_K(X)$  as predictors

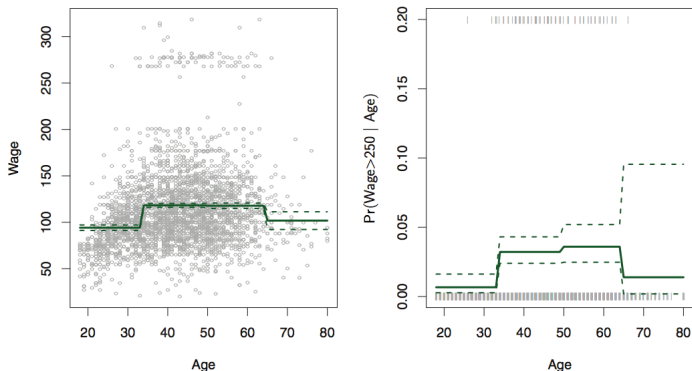
$$y_i = \beta_0 + \beta_1 C_1(x_i) + \beta_2 C_2(x_i) + \dots + \beta_K C_K(x_i) + \epsilon_i,$$

where  $\epsilon_i$  is the error term. (Why there is no  $C_0(X)$  in the model?)

- $\beta_j$  represents the average increase in the response for  $X$  in  $c_j \leq X < c_{j+1}$  relative to  $X < c_1$ .

# The Wage Data

## Piecewise Constant



Left: The solid blue curve is a step function of wage as a function of age, fit by least squares. The dotted curves indicate an estimated 95 % confidence interval. Right: We model the binary event  $\text{wage} > 250$  using logistic regression, with the step function.

# Pros and Cons of Step Function

- The step function approach is widely used in biostatistics and epidemiology among other areas, because the model is easy to fit and the regression coefficient has a natural interpretation.
- However, unless there are natural breakpoints in the predictors, piecewise-constant functions can miss the trend of the curve. The choice of breakpoints can be problematic.
- Polynomial and piecewise-constant regression models are in fact special cases of a **basis function** approach,

$$y_i = \beta_0 + \beta_1 b_1(x_i) + \beta_2 b_2(x_i) + \dots + \beta_K b_K(x_i) + \epsilon_i,$$

where  $b_1(X), b_2(X), \dots, b_K(X)$  are known basis functions.

- In the following, we investigate a very common choice for a basis function: **regression splines**.



# Piecewise Polynomials

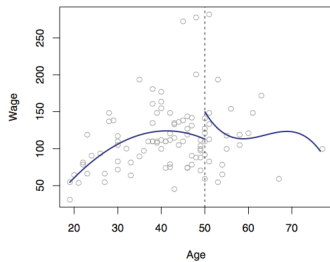
- Instead of a single polynomial in  $X$  over its whole domain, we can rather use different polynomials in regions defined by knots,

$$y_i = \begin{cases} \beta_{01} + \beta_{11}x_i + \beta_{21}x_i^2 + \beta_{31}x_i^3 + \epsilon_i & \text{if } x_i < c; \\ \beta_{02} + \beta_{12}x_i + \beta_{22}x_i^2 + \beta_{32}x_i^3 + \epsilon_i & \text{if } x_i \geq c. \end{cases}$$

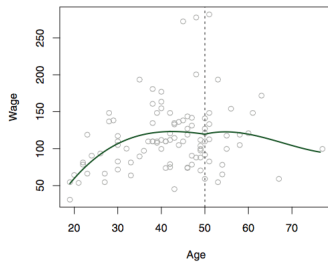
- Using more knots leads to a more flexible piecewise polynomial. In general, if we place  $K$  different knots throughout the range of  $X$ , then we will end up fitting  $K + 1$  different cubic polynomials.
- Better to add constraints to the polynomials, e.g. continuity. This leads to **cubic splines**.
- The general definition of a degree- $d$  spline is that it is a piecewise degree- $d$  polynomial, with continuity in derivatives up to degree  $d - 1$  at each knot.

# The Wage Data

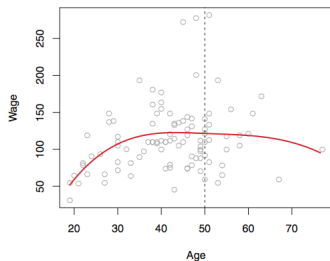
### Piecewise Cubic



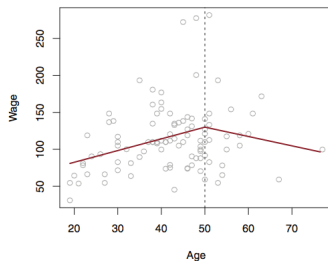
### Continuous Piecewise Cubic



### Cubic Spline



### Linear Spline



# The Spline Basis Representation

- How can we construct the degree-d spline?
- A **linear spline** with knots at  $\xi_k, k = 1, \dots, K$  is a piecewise linear polynomial continuous at each knot. It is

$$y_i = \beta_0 + \beta_1 b_1(x_i) + \beta_2 b_2(x_i) + \dots + \beta_{K+1} b_{K+1}(x_i) + \epsilon_i,$$

where  $b_k$  are basis functions

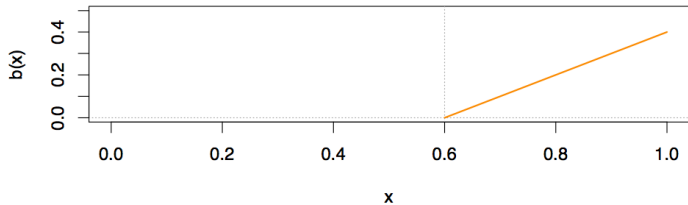
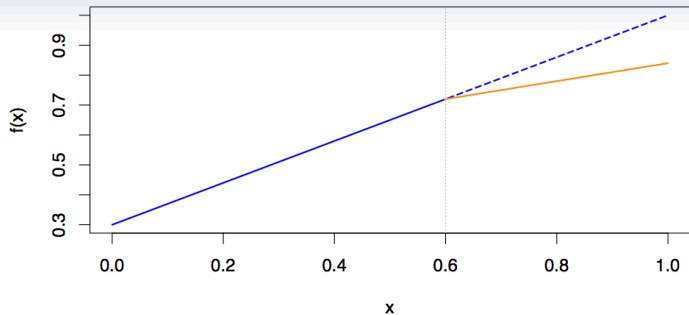
$$b_1(x_i) = x_i, b_{k+1}(x_i) = (x_i - \xi_k)_+, \quad k = 1, \dots, K,$$

here  $(\cdot)_+$  means positive part,

$$(x_i - \xi_k)_+ = \begin{cases} x_i - \xi_k & \text{if } x_i > \xi_k \\ 0 & \text{otherwise} \end{cases}$$

- What is the interpretation of  $\beta_1$ ? (The averaged increase of  $Y$  if we increase one unit of  $X$  when  $X < \xi_1$ .)

# Linear Splines



- A **cubic spline** with knots at  $\xi_k, k = 1, \dots, K$  is a piecewise cubic polynomial with continuous derivatives up to order 2 at each knot. It is

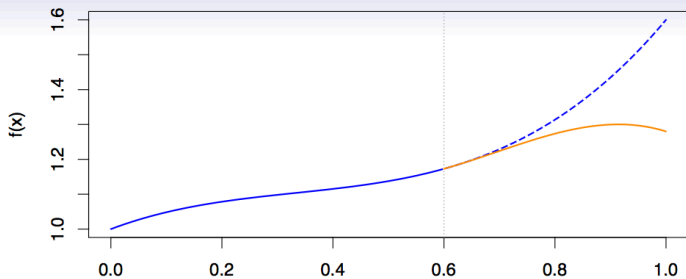
$$y_i = \beta_0 + \beta_1 b_1(x_i) + \beta_2 b_2(x_i) + \dots + \beta_{K+3} b_{K+3}(x_i) + \epsilon_i,$$

where  $b_k$  are basis functions

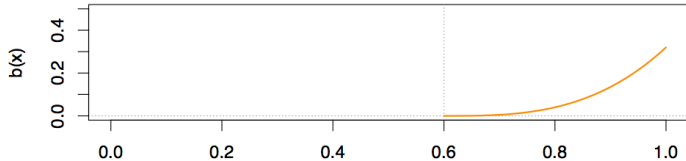
$$b_1(x_i) = x_i, b_2(x_i) = x_i^2, b_3(x_i) = x_i^3,$$

$$b_{k+3}(x_i) = (x_i - \xi_k)_+^3, \quad k = 1, \dots, K,$$

# Cubic Splines



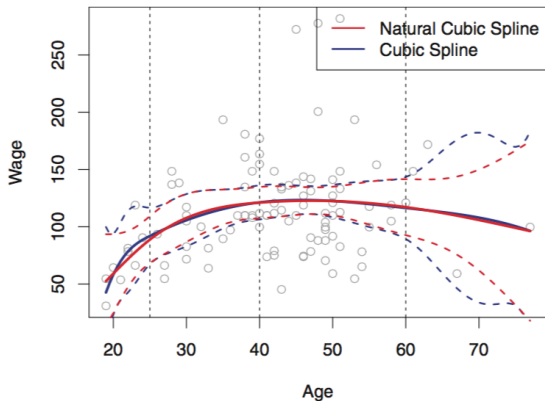
x



x

# Natural Cubic Splines

A natural spline is a regression spline with additional boundary constraints: the function is required to be linear at the boundary.



# Choosing the Number and Locations of the Knots

- Typically, we place  $K$  knots at the corresponding quantiles of the data or place on the range of  $X$  with equal space. Usually, the placement of knots is not very crucial.
- We use cross-validation to choose  $K$ . Specifically, given a fixed  $K$ , we use cross-validation to estimate the test RSS, and then we choose  $K$  with smallest estimated test RSS.



# Smoothing Splines

- The **smoothing spline** is the minimizer of the following objective function

$$\sum_{i=1}^n (y_i - g(x_i))^2 + \lambda \int g''(t)^2 dt,$$

where  $\lambda$  is a nonnegative tuning parameter.

- The first term  $\sum_{i=1}^n (y_i - g(x_i))^2$  is RSS which tries to make  $g(x)$  match the data at each  $x_i$ .
- Broadly speaking, the second derivative of a function is a measure of its roughness. So the second term  $\int g''(t)^2 dt$  is a roughness penalty on the entire range of  $X$ .
- The tuning parameter  $\lambda$  determines the importance between the model fit and the smoothness of the estimated function (bias-variance trade-off).
- It can be shown that the minimizer is a shrunken version of the natural cubic spline with knots at  $x_1, \dots, x_n$ . The math is beyond this lecture, and we will not pursue this approach.

# Local Regression

**Local regression** is a different approach for fitting flexible non-linear functions, which involves computing the fit at a target point  $x_0$  using only the nearby training observations.

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**Algorithm 7.1** *Local Regression At  $X = x_0$* 

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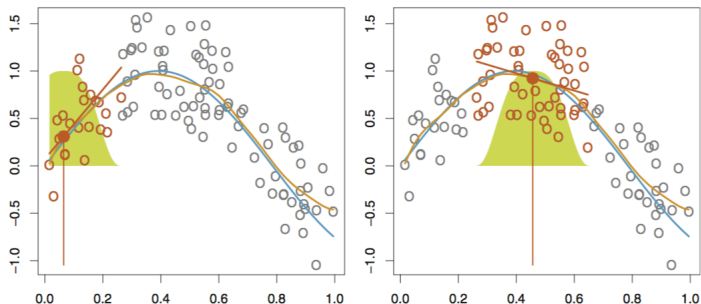
1. Gather the fraction  $s = k/n$  of training points whose  $x_i$  are closest to  $x_0$ .
2. Assign a weight  $K_{i0} = K(x_i, x_0)$  to each point in this neighborhood, so that the point furthest from  $x_0$  has weight zero, and the closest has the highest weight. All but these  $k$  nearest neighbors get weight zero.
3. Fit a *weighted least squares regression* of the  $y_i$  on the  $x_i$  using the aforementioned weights, by finding  $\hat{\beta}_0$  and  $\hat{\beta}_1$  that minimize

$$\sum_{i=1}^n K_{i0} (y_i - \beta_0 - \beta_1 x_i)^2. \quad (7.14)$$

4. The fitted value at  $x_0$  is given by  $\hat{f}(x_0) = \hat{\beta}_0 + \hat{\beta}_1 x_0$ .

# Simulated Example

## Local Regression



The blue curve is true  $f(x)$ , and the light orange curve is the local regression  $\hat{f}(x)$ . The orange colored points are local to the target point  $x_0$ , represented by the orange vertical line. The yellow bell-shape indicates weights assigned to each point, decreasing to zero with distance from the target point. The fit  $\hat{f}(x_0)$  at  $x_0$  is obtained by fitting a weighted linear regression (orange line segment), and using the fitted value at  $x_0$  (orange solid dot) as the estimate  $\hat{f}(x_0)$ .

# Local Regression

- The size of the neighborhood (fraction  $s$  of training data) is a tuning parameter, which can be chosen by cross-validation.
- When we have two dimensional predictors  $X_1$  and  $X_2$ , we can simply use two-dimensional neighborhoods, and fit bivariate linear regression models using the observations that are near each target point in two-dimensional space.
- However, local regression can perform poorly if  $p$  is much larger than about 3 or 4 (known as curse of dimensionality).

# Generalized Additive Models

- In the previous sections, we only have a single predictor  $X$ .
- **Generalized additive models** (GAMs) provide a general framework for extending a standard linear model by allowing non-linear functions of each of the variables, while maintaining additivity,

$$y_i = \beta_0 + f_1(x_{i1}) + f_2(x_{i2}) + \dots + f_p(x_{ip}) + \epsilon_i.$$

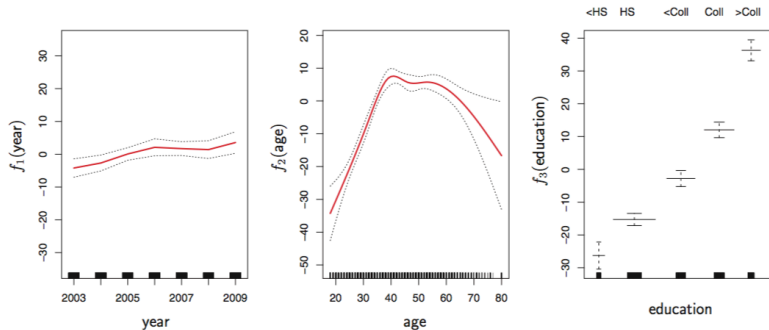
- We can fit GAMs using smoothing splines or other smoothing methods (local regression, regression splines) for a single predictor, via an approach known as backfitting.
- Coefficients not that interesting; fitted functions are.
- Can mix terms – some linear, some nonlinear.
- Can be applied to classification problems

$$\text{logit } P(y_i = 1|x_i) = \beta_0 + f_1(x_{i1}) + f_2(x_{i2}) + \dots + f_p(x_{ip}).$$

# Wage Data

Consider the wage data

$$wage = \beta_0 + f_1(year) + f_2(age) + f_3(education) + \epsilon.$$

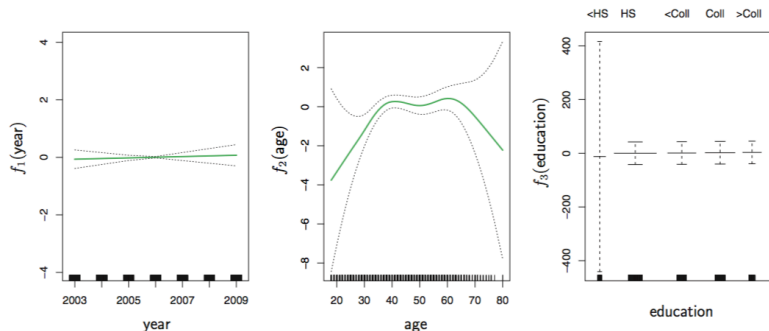


The first two functions are natural splines in year and age, with four and five degrees of freedom, respectively. The third function is a step function, fit to the qualitative variable education.

# Wage Data

Consider the wage data

$$\text{logit } P(\text{wage} > 250) = \beta_0 + \beta_1 \times \text{year} + f_2(\text{age}) + f_3(\text{education}).$$



The first function is linear in year, the second function a smoothing spline with five degrees of freedom in age, and the third a step function for education. There are very wide standard errors for the first level <HS of education.

# Pros and Cons of GAMs

- GAMs allow us to fit a non-linear  $f_j$  to each  $X_j$ , so that we can automatically model non-linear relationships that standard linear regression will miss.
- The non-linear fits can potentially make more accurate predictions for the response  $Y$ .
- Because the model is additive, we can still examine the effect of each  $X_j$  on  $Y$  individually while holding all of the other variables fixed.
- It solves the curse of dimensionality.
- However, GAMs fail to incorporate the interaction of variables.