

Matrix-valued Time Series in High Dimension

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 - Introduction: model and objective
 - Procedures and theoretical guarantees
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- $\Sigma \in \mathcal{S}_p^{++}$ has a Toeplitz structure .

$$\Sigma := \begin{pmatrix} \sigma_0 & \sigma_1 & \sigma_2 & \sigma_3 & \sigma_4 & \cdots & \sigma_{p-1} \\ \sigma_1 & \sigma_0 & \sigma_1 & \sigma_2 & \sigma_3 & \cdots & \sigma_{p-2} \\ \vdots & \ddots & \ddots & \ddots & \ddots & \ddots & \vdots \\ \vdots & \ddots & \sigma_1 & \sigma_0 & \sigma_1 & \ddots & \vdots \\ \vdots & \ddots & \ddots & \ddots & \ddots & \ddots & \vdots \\ \sigma_{p-2} & \cdots & \sigma_3 & \sigma_2 & \sigma_1 & \sigma_0 & \sigma_1 \\ \sigma_{p-1} & \cdots & \sigma_4 & \sigma_3 & \sigma_2 & \sigma_1 & \sigma_0 \end{pmatrix}$$

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- We also develop a procedure that selects non-null correlation coefficients.
- Numerical results illustrate the excellent behaviour of the test procedures and the support selector.

- The one-sided test problem is

$$H_0 : \Sigma = I_p, \quad \text{vs.} \quad H_1 : \Sigma \in \mathcal{F}_+(s, S, \sigma),$$

where

$$\mathcal{F}_+(s, S, \sigma) = \left\{ \Sigma \in \mathcal{S}_p^{++} \cap \mathcal{T}_p \text{ and } \exists \mathcal{C} \subseteq \{1, \dots, S\}, \right. \\ \left. \begin{array}{ll} |\mathcal{C}| = s, \forall j \in \{1, p-1\}, & \sigma_j \geq \sigma > 0, & j \in \mathcal{C}, \\ & \sigma_j = 0, & j \notin \mathcal{C} \end{array} \right\}$$

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- The two-sided test problem is

$$H_0 : \Sigma = I_p, \quad \text{vs. } H_1 : \Sigma \in \mathcal{F}(s, S, \sigma),$$

where $\mathcal{F}(s, S, \sigma)$ is defined similarly as $\mathcal{F}_+(s, S, \sigma)$ by considering the absolute values of the covariance elements.

Moderately sparse case in the one-sided alternative

- When the alternative is $\mathcal{F}_+(s, S, \sigma)$, we consider for some threshold $t_{n,p}^{MS+}$ the test procedure

$$\Delta_n^{MS+} = \mathbb{1} \left(\text{Sum}_{\{1:S\}}^+(\Sigma_n - I_p) \geq t_{n,p}^{MS+} \right),$$

where for an arbitrary set $\mathcal{C} \subseteq \{1, \dots, S\}$,

$$\text{Sum}_{\mathcal{C}}^+(\Sigma_n) := \sum_{j \in \mathcal{C}} \text{Tr}(A_j \Sigma_n) = \sum_{j \in \mathcal{C}} \hat{\sigma}_j.$$

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- When the alternative is $\mathcal{F}(s, S, \sigma)$, we consider for some threshold $t_{n,p}^{MS}$ a test Δ_n^{MS} that sums the absolute values of the first S covariance elements of $\Sigma_n - I_p$ and compare it to $t_{n,p}^{MS}$.

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Theorem (B., Butucea, Sorba 2022)

For $u > 0$, consider $t_{n,p}^{MS+} = \max \left\{ \sqrt{\frac{u \cdot S}{n(p-S)}}, \frac{2u \cdot S}{n(p-S)} \right\}$. Then $R(\Delta_n^{MS+}, \mathcal{F}_+) \leq 2 \exp(-\frac{u}{4})$ provided that $\sigma \geq \frac{2(s+1)}{s} t_{n,p}^{MS+}$.

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Theorem (B., Butucea, Sorba 2022)

For $u > 1$, consider $t_{n,p}^{HS+} = \max \left\{ \sqrt{\frac{4u \cdot s \log \binom{S}{s}}{n(p-S)}}, \frac{8u \cdot s \log \binom{S}{s}}{n(p-S)} \right\}$. Then

$R(\Delta_n^{HS+}, \mathcal{F}^+) \leq \exp \left(-(u-1) \log \binom{S}{s} \right) + \exp \left(-\frac{u}{4} \right)$ provided that

$\sigma \geq \frac{1}{s} \left(t_{n,p}^{HS+} + (2s+1) \max \left\{ \sqrt{\frac{u \cdot s}{n(p-S)}}, \frac{2u \cdot s}{n(p-S)} \right\} \right)$.

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Different structured matrix estimation is studied in Klopp, Lu, Tsybakov, Zhou 2019 (Bernoulli)

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- Global idea: $Y \longrightarrow Y_r \longrightarrow \hat{A}_r X \hat{B}_r$.
- **Identifiability:** The predictors are not uniquely defined in this setting. Without further strong assumptions, we cannot hope to learn parameters from a non identifiable model.

Diagonal 2MR

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One to one mapping between A^*/A_0^* and B^*/B_0^* . E_0 and E share the same singular values.

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The initial problem is equivalent to finding predictors satisfying

$$(\hat{A}_{0r}, \hat{B}_{0r}) \in \underset{\substack{A_0, B_0: \\ \text{rank } A_0 \wedge \text{rank } B_0 \leq r}}{\arg \min} \|\Sigma_Y - A_0 \Sigma_X B_0\|_F^2.$$

Solution of D2MR

- **Objective:** Under the constraint $\text{rank}(A_0) \leq r$ and $\text{rank}(B_0) \leq r$, minimize:

$$\left\| \underbrace{\begin{pmatrix} \sigma_1(Y) & & & \\ & \ddots & & \\ & & \sigma_{r_Y}(Y) & \\ & & & 0 \end{pmatrix}}_{n \times p} - A_0 \underbrace{\begin{pmatrix} \sigma_1(X) & & & \\ & \ddots & & \\ & & \sigma_{r_X}(X) & \\ & & & 0 \end{pmatrix}}_{m \times q} B_0 \right\|_F^2 .$$

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- **Solution:**

$$\hat{A}_{0r} = \underbrace{\begin{pmatrix} \sigma_1(Y) & & & \\ & \ddots & & \\ & & \sigma_{r \wedge r_Y}(Y) & \\ & & & 0 \end{pmatrix}}_{n \times m}, \quad \hat{B}_{0r} = \underbrace{\begin{pmatrix} \sigma_1(X)^{-1} & & & \\ & \ddots & & \\ & & \sigma_r(X)^{-1} & \\ & & & 0 \end{pmatrix}}_{q \times p}.$$

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- How far is the predictor $\hat{A}_{0r} \Sigma_X \hat{B}_{0r}$ from the signal $A^* X B^*$?

Oracle inequality in the fixed rank case

Theorem (B., Butucea 2023)

The predictor $\hat{A}_{0,r}\Sigma_X\hat{B}_{0,r}$ satisfies for $C > 0$ and for any $t > 0$:

$$\|A_0^*\Sigma_X B_0^* - \hat{A}_{0,r}\Sigma_X\hat{B}_{0,r}\|_F^2 \leq 9 \inf_{\substack{A_0, B_0: \\ \text{rank } A_0 \wedge \text{rank } B_0 \leq r}} \|A_0^*\Sigma_X B_0^* - A_0\Sigma_X B_0\|_F^2 + C\sigma^2(1+t)^2 \cdot r(n+p),$$

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- $\mathcal{O}(r(n+p))$ is the minimax optimal rate in the (one-sided) *matrix regression* (MR) model.
- From the explicit solutions $(\hat{A}_{0r}, \hat{B}_{0r})$ we deduce (\hat{A}_r, \hat{B}_r) solution to the initial problem:

$$\hat{A}_r = U_Y \hat{A}_{0r} U_X^\top \quad \text{and} \quad \hat{B}_r = V_X \hat{B}_{0r} V_Y^\top.$$

Further results

We derive a rank-adaptive procedure.

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Simulation results confirm the good prediction and the rank consistency results under data-driven explicit choices of the tuning parameters and the scaling parameter of the noise.

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- 1 Covariance matrix testing and support recovery: Chap. 2
- 2 Two-Sided Matrix Regression: Chap. 3
- 3 Dynamic Topic Model: Chap. 4 & 5
 - Introduction: Topic Models, Identifiability, Dynamic.
 - Dynamic Latent Factors: Procedure and theoretical guarantees
 - Dynamic Topic Model: Procedure and theoretical guarantees

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$$\Pi^* = A^* W^*,$$

where $A^* \in \mathbb{R}^{p \times K}$ has columns in \mathcal{S}_{p-1} , $W^* \in \mathbb{R}^{K \times n}$ has columns in \mathcal{S}_{K-1} .

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- **Interpretation:**

$$\mathbb{P}(\text{word } i | \text{document } j) = \sum_{k=1}^K \mathbb{P}(\text{word } i | \text{topic } k) \mathbb{P}(\text{topic } k | \text{document } j)$$

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Bibliography

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Double randomness: Dirichlet + Multinomial

- Let the **topic-topic overlapping matrix** measure the affinity of topics using the same words:

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Assume: $\lambda_K(\Sigma_A) \geq c$, $\min_{k,l} [\Sigma_A]_{kl} \geq c$ and $\min_i h_i := h_{\min} \geq c \frac{K}{p}$.

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$$\Sigma_W^{1:T} := \frac{1}{nT} \left(W^{1:T} \right) \left(W^{1:T} \right)^\top,$$

capture the affinity of topics to be covered together in the same document.

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Assume: $\lambda_K(\Sigma_{\mathbf{W}}^{1:T}) \geq c > 0$, a.s..

Remark: if $\min_k \tilde{\theta}_k^* \geq c > 0$, this holds for large enough n , T with high probability.

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$$\min_{k \in [K]} \tilde{\theta}^*(k) \geq \underline{\theta} \text{ and } m \leq \text{Tr}(\Sigma(\theta^*)) \leq 1.$$

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$$G_\eta := \left\{ x : x = \sum_{k=1}^K \alpha_k \boldsymbol{\eta}_k, \forall k \in [K], \alpha_k \geq 0, \sum_{k=1}^K \alpha_k = 1 \right\}.$$

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Define $\Lambda \in \mathbb{R}^{p \times K}$ by solving for all $i \in [p]$,

$$[R]_i = \sum_{k=1}^K [\Lambda]_{ik} \eta_k,$$

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- *Word-topic matrix estimation*: Define $\Gamma := M_*^{1/2} \text{diag}([U]_{\cdot 1}) \Lambda$. Normalize each column of Γ by its \mathbb{L}_1 norm. The resulting matrix is A^* .

Dynamic Latent Factors: Estimators

- We define $\hat{\theta}$, estimator of $\tilde{\theta}^*$, as the empirical mean of the recovered $(W_j^{t+1})_{j,t}$:

$$\hat{\theta} := \frac{1}{n(T-1)} \sum_{j=1}^n \sum_{t=1}^{T-1} W_j^t.$$

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- We estimate $1 - c^*$ by the normalized sum of scalar products:

$$\widehat{(1-c)} := \frac{\sum_{t=1}^{T-1} \sum_{j=1}^n \langle W_j^{t+1} - \overline{W}^{+1}; W_j^t - \overline{W} \rangle}{\sum_{t=1}^{T-1} \sum_{j=1}^n \|W_j^t - \overline{W}\|_2^2},$$

$$\overline{W}^{+1} := \frac{1}{n(T-1)} \sum_{t=1}^{T-1} \sum_{j=1}^n W_j^{t+1} \text{ and } \overline{W} := \frac{1}{n(T-1)} \sum_{t=1}^{T-1} \sum_{j=1}^n W_j^t.$$

Using the variance of the stationary sequence and the explicit expression of the matrix Σ , we see that:

$$\text{Tr}(\mathbb{V}(w_j^t)) = \frac{c^*}{2 - c^*} \frac{1 - \|\tilde{\theta}^*\|_2^2}{\alpha + 1}.$$

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$$\hat{\alpha} = \frac{\hat{c}}{2 - \hat{c}} \frac{1 - \|\hat{\theta}\|_2^2}{\mathcal{V}} - 1, \quad \text{where} \quad \mathcal{V} := \frac{1}{n(T-1)} \sum_{t=1}^{T-1} \sum_{j=1}^n \|w_j^t - \bar{w}\|_2^2.$$

Theorem (B., Butucea, Ke 2024)

For any N , n and T large enough, with probability at least $1 - \frac{C_1}{nT}$:

$$\max \left\{ \left\| \hat{\theta} - \tilde{\theta}^* \right\|_2, \left| \widehat{(1-c)} - (1-c^*) \right|, \left| \hat{\alpha} - \alpha^* \right| \right\} \leq C_2 \cdot \sqrt{\frac{\log(nT)}{n(T-1)}},$$

where $C_1, C_2 > 0$ are explicit constants, free of the dimensions appearing in the model.

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- 3 *Behaviour of the vertex hunting algorithm with noisy entries.*

Theorem (B., Butucea, Ke 2024)

For N , n and T large enough, there exists χ , a positive constant only depending on K , such that with probability at least $1 - \frac{8}{nT}$:

$$\sum_{i=1}^p \left\| [\hat{A}]_{i.} - [A^*]_{i.} \right\|_1 \leq \chi \sqrt{\frac{p \log(nT) + p^2}{nT(N-2)}} p(1+p) \left(1 + \max_{x \in \mathcal{G}_n} \|x\|_2\right).$$

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For N , n and T large enough, and fixed number of topics K and of the vocabulary size p , with probability at least $1 - \frac{C}{nT}$:

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The bounds are driven by the Dirichlet noise and by the multinomial noise.